Exposure to arsenic has long been known to have neurologic consequences in adults, but to date there are no well-controlled studies in children. We report results of a cross-sectional investigation of intellectual function in 201 children 10 years of age whose parents participate in our ongoing prospective cohort study examining health effects of As exposure in 12,000 residents of Araihazar, Bangladesh. Water As and manganese concentrations of tube wells at each child’s home were obtained by surveying all wells in the study region. Children and mothers came to our field clinic, where children received a medical examination in which weight, height, and head circumference were measured. Children’s intellectual function on tests drawn from the Wechsler Intelligence Scale for Children, version III, was assessed by summing weighted items across domains to create Verbal, Performance, and Full-Scale raw scores. Children provided urine specimens for measuring urinary As and creatinine and were asked to provide blood samples for measuring blood lead and hemoglobin concentrations. Exposure to As from drinking water was associated with reduced intellectual function after adjustment for sociodemographic covariates and water Mn. Water As was associated with reduced intellectual function, in a dose–response manner, such that children with water As levels > 50 µg/L achieved significantly lower Performance and Full-Scale raw scores than did children with water As levels ≤ 5.5 µg/L. The association was generally stronger for well-water As than for urinary As. Key words: arsenic, children, IQ. Environ Health Perspect 112:1329–1333 (2004). doi:10.1289/ehp.6964 available via http://dx.doi.org/
school or was > 10.5 years of age (n = 33), the family was not at home, or other, unspecified reasons (n = 64). To bolster our sample size, in 25 instances when a home visit identified an excluded child, interviewers selected a child of participating parents from the same village until 201 children, using 196 wells, had been assessed.

Procedure. Children and their mothers came to our field clinic, where children participated in the assessments described below and received a medical examination by a study physician. Weight, height, and head circumference were measured. Children provided urine specimens for the measurement of urinary As and creatinine and were asked to provide a blood sample for the measurement of blood lead (BiPb) and hemoglobin (Hgb) concentrations. Of the 201 children assessed, 107 agreed to provide blood samples. Information on family demographics (e.g., parental education, occupation, housing type) was available from the baseline interview of parents during their enrollment in the cohort study. Information on the primary source of drinking water was obtained from the child’s mother. Parents were asked whether their home included a television; about parental age, education, and occupation; and about child birth order. For an additional surrogate for social class, the type of roofing on the well owner’s home was recorded as thatched, tin, or cement (thatched lowest, cement highest). Children were given a toy as thanks for their participation; families participating in the larger cohort study receive primary medical care at our own field clinic.

Measures. Water analyses. Water As concentrations of tube wells at each child’s home were obtained during a survey of all wells in the study region (van Geen et al. 2003b) and shipped to Columbia University’s Lamont Doherty Earth Observatory for analysis. Water samples were analyzed by graphite furnace atomic absorption (GFAA), which had a detection limit of 5 µg/L. Those water samples found to have < 5 µg/L were subsequently reanalyzed by inductively coupled plasma–mass spectrometry (ICP-MS), which has a detection limit of 0.1 µg/L (Cheng et al. 2004). Of the 196 well-water samples, 194 were also analyzed for Mn by standard flame atomic absorption spectrophotometry.

Biochemical measurements. Urinary As concentrations were assayed by GFAA at the Mailman School of Public Health, using a Perkin-Elmer Analyst 600 system as previously described (Nixon et al. 1991). Our laboratory participates in a quality control program coordinated by P. Weber at the Quebec Toxicology Center (Quebec City, Quebec, Canada). During the course of this study, intraassay correlation coefficients between our laboratory’s values and samples calibrated at Weber’s laboratory were 0.99. Levels of As in urine were also adjusted for urinary creatinine levels, which were analyzed by a colorimetric Sigma Diagnostics Kit (Sigma, St. Louis, MO, USA). In addition, urinary As metabolites were specified using a method adapted after Heitkemper et al. (2001). This method employs high-performance liquid chromatography separation of arsenobetaine (AsB), arsenocholine (AsC), arsenate, arsenite, monomethylarsonic acid (MMA), and dimethylarsinic acid (DMA), followed by detection by ICP-MS. The percentages of inorganic As (InAs; i.e., arsenate + arsenite), MMA, and DMA were calculated after subtracting AsC and AsB from total urinary As.

Venous blood samples were obtained for measurements of BiPb (Fernandez and Hilligoss 1982) and Hgb. Whole-blood samples were appropriately stored and transported to a laboratory at Columbia University that participates in the BiPb quality control program of the Centers for Disease Control and Prevention (CDC; Atlanta, GA, USA). Intraassay correlation coefficients between our laboratory’s values and samples calibrated at CDC ranged between 0.97 and 0.99. Children providing blood samples had mothers with significantly more years of education and received significantly higher Verbal raw scores than did mothers of children not providing blood samples (Wilcoxon tests, degrees of freedom = 1, p-values < 0.05); there were no other differences between those providing and not providing blood samples.

Children’s intellectual function. The Wechsler Intelligence Scale for Children, version III (WISC-III; Wechsler 1991), suitable for children ≥ 6 years of age, consists of five (or six, depending upon administration) verbal subtests, which together provide a Verbal intellectual quotient (IQ) score, and a similar number of performance subtests that together provide a Performance IQ. Neither the WISC-III (Wechsler 1991) nor any other recently well-standardized child IQ test has been adapted or standardized for use in Bangladesh.

In Araihazar, living conditions differ dramatically from those in Western settings where this test was developed, which necessitated adaptations for use in this culture. For example, a typical home consists of a single room, often with a dirt floor. Most families use biomass fuel (leaves, hay, dung) for cooking. Electricity is available in most homes; commonly this consists of one or two bulbs used for lighting. Many common Western household items, such as telephones and bathtubs, are rare.

We used six subtests that seemed the most culturally adaptable to this cultural context. Of the WISC-III Verbal subtests, we used Similarities and Digit Span: Of the Performance subtests, we used Picture Completion, Coding, Block Design, and Mazes. Two items with no recognizable analog were eliminated from the Picture Completion subscale (telephone, bathtub), and close substitutions were made for four others from the Similarities subscale (“mango and banana” for “apple and banana”; “flute and drum” for “piano and guitar”; “dog and cow” for “cat and mouse”; and “tire and ball” for “wheel and ball”). The WISC-III subtests include items of graduated difficulty, with more points awarded for harder items or faster completion. We summed these weighted items across Verbal, Performance, and Full-Scale domains to create Verbal, Performance, and Full-Scale raw scores; we also transformed these into measures of estimated Verbal, Performance, and Full-Scale IQ, using procedures presented in the test manual (Wechsler 1991), despite the obvious limitations in application to this population. Below we use “IQ” to represent this estimated measure.

Maternal intelligence was assessed with Raven’s Standard Progressive Matrices, a nonverbal test relatively free of cultural influences (Raven et al. 1983).

Translation and training. All tests and interviews were translated (and back-translated) between Bangla (Bengali) and English. As noted, items deemed to be culturally inappropriate were altered or omitted. Materials were piloted to ensure maternal and child comprehension; two interviewers were then trained by a competent tester (G.A.W.) and then continued with supervised practice sessions for 2 weeks. All written test responses were rechecked when data were sent to the Columbia University Department of Psychiatry for entry.

Statistical analyses. Outcomes. Because of concerns about the application of U.S. standardization of the WISC-III to Bangladeshi children, we first conducted analyses that predicted Verbal, Performance, and Full-Scale raw scores. Because the psychometric properties of IQ scores are more familiar to readers, we also applied the same analytic models to the prediction of estimated Verbal “IQ,” Performance “IQ,” and Full-Scale “IQ.”

Covariate adjustment. In Bangladesh, grammar school extends to fifth grade. Therefore, mother’s and father’s education were categorized as none, 1–5 years, and 6–13 years. Parental occupation was recoded as laborer/farmer, factory/other paid job, business, or missing/other. Because just 6% (n = 11) of mothers reported working outside the home, only paternal occupation was included in the regression models. From a pool of potential demographic covariates, we retained those that were empirically or theoretically importantly related to child intelligence, as well as those that made an initial contribution (at significance p ≤ 0.20 or better), in initial regression analyses, either to any of the outcomes of interest or to the measures of As exposure.
Analytic model. The analyses first sought to predict the outcomes of interest from the set of sociodemographic factors using linear regression models; once this “core” model was derived, we examined the incremental association of exposures (Mn, As) singly and together, measured continuously. We repeated our analyses, categorizing children into groups, based on quartiles of water As to illustrate dose–response relationships. We next repeated these analyses for the subset of children providing blood samples for the measurement of BPb and Hgb, measured continuously. In all analyses, BPb and water As were log-transformed and water Mn was square root–transformed to make distributions approximately symmetric.

For the most part, analyses are based on n = 201 children. Analyses considering well-water Mn employ n = 194. Analyses involving urinary As and its metabolites are based on n = 200; for analyses considering BPb and Hgb, n = 107.

Results

Sample characteristics. Table 1 presents descriptive information for all demographic, water, and biochemical variables. Average child age was 10 years; approximately half the children in the sample were male; one-third had regular access to a television. On average, mothers and fathers reported 3.7 and 2.9 years of education, respectively. Average child height was 125.6 cm, and average weight was 21.9 kg, values that correspond to roughly the fourth percentile by U.S. norms (CDC 2003b).

Exposure characteristics. Water As concentrations ranged from 0.094 to 790 µg/L, with a mean (117.8 µg/dL) and distribution comparable to those in the larger set of 6,000 contiguous wells in Araihazar (van Geen et al. 2003b). The mean water Mn concentration of 1.386 µg/L was well in excess of the U.S. and WHO recommended maximum contaminant level (MCL) of 500 µg/L, with a range up to 5,438; water Mn values are not available for the entire set of 6,000 wells. Indeed, 82% of children were consuming water in excess of the MCL for Mn. The association between water As and water Mn was significant (Spearman \( r = 0.39; p < 0.0001 \)) but not strong enough to preclude examination of their independent effects on child intelligence. The correlation between water As and urinary As (Spearman \( r = 0.45; p < 0.0001 \)) was comparable to that previously reported for adults in this region (Ahsan et al. 2000).

The subset of children for whom BPb measures were obtained (n = 107). Spearman correlations (necessitated by skewed distributions) with well-water As (−0.16) and with urinary As (−0.06) were not significant.

Relationship between covariates and intellectual function. Linear regression analyses predicting test raw scores from the sociodemographic features retained in the final “core” model revealed generally better scores in children of more educated mothers and of mothers with higher Raven scores, those living in more adequate dwellings, those with access to television, and those who were taller and had larger head circumference (data not shown).

Relationship between well-water Mn and intellectual function. Without adjustment for either core variables or water As, water Mn was significantly associated with Full-Scale and Performance raw scores (\( B \)-values = −0.33 and −0.29, \( p < 0.002 \) and \( p < 0.001 \), respectively) but not with Verbal raw score (\( B = −0.04, p = 0.15 \)). Addition of Mn to analytic models made little change in associations between core model variables and intellectual function raw scores. Controlling for sociodemographic features, Mn levels were significantly negatively associated with Performance and Full-Scale raw scores (\( B \)-values = −0.20 and −0.22, respectively, \( p < 0.03 \)) but not with Verbal raw score (\( B = −0.02, p > 0.5 \)). However, when water As was added to these models, Mn made no significant (\( p > 0.25 \)) contribution to intellectual function. With both water As and water Mn in the model, there was no significant interaction in their prediction of Full-Scale, Verbal, or Performance raw scores.

Relationship between well-water As and intellectual function. Table 2 presents associations between water As and intellectual function, before and after adjustment for sociodemographic features. In each case, associations between water As and intellectual function raw scores were stronger before adjustment for sociodemographic features. In unadjusted analyses, water As explained 7.29, 2.61, and 7.04% of the variance in Performance, Verbal, and Full-Scale raw scores, respectively. With covariate adjustment, water As remained significantly negatively associated with both performance and Full-Scale raw scores, explaining an incremental 4.33, 0.89, and 3.88% of the variance in Performance, Verbal, and Full-Scale raw scores, respectively. Results were similar when “IQ” outcomes were substituted for raw scores (data not shown).

Dose–response relationships between water As and intellectual function. Figure 1 illustrates the adjusted Full-Scale, Performance, and Verbal raw scores by As quartile. As water As increased, there were dose-dependent changes in adjusted and unadjusted (data not shown) scores. With adjustment, compared with the lowest quartile of As exposure, the third and fourth quartiles had significantly lower scores on both Full-Scale (\( B = −7.8 \) and −11.3, \( p < 0.05 \) and \( p < 0.01 \)), respectively) and Performance raw scores (\( B = −7.3 \) and −9.7, \( p < 0.05 \) and \( p < 0.01 \)), respectively. The highest exposure quartile was marginally lower on Verbal raw score than the lowest exposure quartile (\( B = −1.6, p < 0.10 \)). The relationship between water As (measured continuously) and Full-Scale raw score is illustrated in Figure 2. Water As concentrations of 10 and 50 µg/L

**Table 1.** Sample characteristics [no. (%)] or mean ± SD.

<table>
<thead>
<tr>
<th>Variable</th>
<th>No. (%)</th>
<th>Male</th>
<th>Television access</th>
<th>House type</th>
<th>Thatched roof or poorer</th>
<th>Corrugated tin</th>
<th>Concrete construction</th>
<th>Father’s occupation</th>
<th>Other/missing</th>
<th>Laborer/Farmer</th>
<th>Factory/other paid job</th>
<th>Business</th>
<th>Child age</th>
<th>Full-Scale “IQ”</th>
<th>Verbal “IQ”</th>
<th>Performance “IQ”</th>
<th>Full-Scale raw score</th>
<th>Verbal raw score</th>
<th>Performance raw score</th>
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<tbody>
<tr>
<td>Urinary As (µg/L)</td>
<td>296.6 ± 277.2</td>
<td>98 (48.8)</td>
<td>70 (34.8)</td>
<td>20 (10.0)</td>
<td>149 (74.1)</td>
<td>32 (15.2)</td>
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<td>10.0 ± 0.4</td>
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<td>53.0 ± 6.3</td>
<td>55.4 ± 5.2</td>
<td>58.4 ± 8.0</td>
<td>70.5 ± 20.8</td>
<td>16.5 ± 6.1</td>
<td>54.0 ± 17.4</td>
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<td>Height (cm)</td>
<td>125.6 ± 6.5</td>
<td>21.9 ± 3.3</td>
<td>13.8 ± 1.1</td>
<td>49.5 ± 1.4</td>
<td>23 (11.1)</td>
<td>14 (7.1)</td>
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<td>10.0 ± 0.4</td>
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<td>117.8 ± 49.2</td>
<td>116.6 ± 132.5</td>
<td>116.6 ± 34.5</td>
<td>296.6 ± 277.2</td>
<td>126.1 ± 1.1</td>
<td>10.1 ± 3.3</td>
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<td>Weight (kg)</td>
<td>21.9 ± 3.3</td>
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<td>Body mass index (kg/m²)</td>
<td>13.8 ± 1.1</td>
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<td>Head circumference (cm)</td>
<td>49.5 ± 1.4</td>
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<td>Mother’s education (years)</td>
<td>3.7 ± 3.7</td>
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<td>Father’s education (years)</td>
<td>3.2 ± 6.7</td>
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<td>Mother’s age (years)</td>
<td>14.4 ± 3.5</td>
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<td>Mother’s Raven’s score</td>
<td>117.8 ± 49.2</td>
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<td>Water Mn (µg/L)</td>
<td>116.6 ± 132.5</td>
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<td>Water As (µg/L)</td>
<td>117.8 ± 145.2</td>
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<td>Urinary As (µg/L)</td>
<td>116.6 ± 148.8</td>
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<td>Urinary creatinine (mg/dL)</td>
<td>43.3 ± 34.1</td>
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<td>Urinary As (µg/g creatinine)</td>
<td>296.6 ± 277.2</td>
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<td>Hgb (g/dL)</td>
<td>12.6 ± 1.1</td>
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<td>BPb (µg/dL)</td>
<td>10.1 ± 3.3</td>
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The table shows the relationship between water arsenic and child intellectual function. The data indicate a strong negative correlation between arsenic exposure and cognitive performance, with significant differences observed between high and low exposure groups. The results also suggest that the relationship is dose-dependent, with incremental decreases in cognitive function observed as arsenic concentrations increase. The analysis was adjusted for sociodemographic factors and revealed that arsenic exposure was a significant predictor of intellectual function in this population. Further studies are needed to confirm these findings and to explore potential mechanisms linking arsenic exposure to cognitive outcomes.
were associated with decrements in Full-Scale raw scores of 3.8 and 6.4 points, respectively.

**Relationship between urinary As, As metabolites, and intellectual function.** We examined relationships between total urinary As concentration, as micrograms per gram creatinine, and child intellectual function. After adjustment for core variables, the associations between urinary As and measures of intellectual function were not statistically significant for Full-Scale ($B = -2.9$, $p = 0.09$), Performance ($B = -2.2$, $p = 0.14$), or Verbal scores ($B = -0.7$, $p = 0.11$) but were in the anticipated direction.

Urine samples were analyzed by high-performance liquid chromatography/ICP-MS for the relative amounts of InAs, MMA, and DMA. The percentages of InAs, MMA, and DMA were calculated after subtracting the contribution of AsC and AsB to total As concentration. The mean ± SD AsC and AsB concentrations were 3.9 ± 3.5 µg/g creatinine and 5.0 ± 7.5 µg/g creatinine, respectively. The frequency distributions of InAs, MMA, and DMA are illustrated in Figure 3. There was a wide variability in the extent to which children eliminated As in the dimethylated form. On average, the percentages of urinary As eliminated as InAs, MMA, and DMA were 12.2, 8.9, and 74.1%, respectively. We posited that children who were poor methylators might be particularly adversely affected by As. However, when both DMA and urinary As were included in the core model, DMA failed to make a significant contribution to intellectual function and did not alter the estimates for total urinary As.

**Relationships between BPb, Hgb, and intellectual function.** Analyses predicting intellectual raw scores from other hematologic measures, adjusted for the same demographic features, were conducted for the subset of 107 children providing blood samples. No significant associations were detected for log BPb or for Hgb on Verbal, Performance, or Full-Scale raw scores or “IQ,” with or without the inclusion of water As (data not shown).

**Discussion**

This is the first systematic study of effects of As on children’s intellectual function. Exposure to As from drinking water was associated with reduced scores on measures of intellectual function, before and after adjusting for sociodemographic features known to contribute to intellectual function. With covariate adjustment, water As remained significantly negatively associated with both Performance and Full-Scale raw scores. Exposure to As was associated with reduced intellectual function, in a dose–response manner, such that children with exposures > 50 µg/L received significantly lower Performance and Full-Scale scores than children with exposures < 5.5 µg/L. The association was stronger for well-water As than for urinary As. Children in the highest quartile of water As scored approximately 10 points lower in Performance raw scores than did those in the lowest quartile.

We have made diligent efforts to reduce the consumption of As-contaminated water in the Araihazar population since our original well survey was conducted in the first half of 2000. For example, each well was labeled to indicate As concentrations > and < 50 µg/L, with either a skull and cross-bones or a picture of a child drinking water. In addition, a village education program that encouraged well switching (van Geen et al. 2002) successfully reached roughly half of all residents. Beyond these, new low-As private and community wells have been installed in parts of the region during this time frame (van Geen et al. 2003a). It is therefore likely that some recent reduction in these children’s As exposure occurred between January 2001 (when well labeling began) and the summer of 2002 (when children were assessed). Indeed, in our simultaneous prospective cohort study in adults, repeated measurements of urinary As concentrations over the same interval have declined. Because urinary As reflects recent exposure, reduced exposure may explain the weaker associations between intellectual function and levels of As in urine, compared with levels in water.

Two published studies of As exposure also found adverse associations with children’s intellectual function. In a small ($n = 80$) sample of children from a Pb smelter area in Mexico, Calderon et al. (2001) found negative associations between children’s urinary As and Verbal intelligence, controlling for a small set of demographic factors. Although investigating anthropogenic exposure to As and Pb, that study did not consider other potential toxins to which nearby residents were exposed. In a second ecological study, Tsai et al. (2003) compared adolescents in Taiwan from regions with and without elevated As in well water, with no measure of individual exposure. With minimal control for sociodemographic factors, adolescents in the exposed group showed inconsistently poorer scores on Performance-type tests; some outcomes were adversely affected in adolescents with low exposure (but not in those with high exposure) relative to those without exposure.

**As metabolism.** Humans excrete MMA and DMA after ingestion of arsenate or arsenite, but the extent of metabolism is remarkably variable and may influence both pre- and postnatal toxicity (Hopenhayn-Rich et al. 1996). This variability in methylation is likely due to both genetic (Chung et al. 2002; Vahic et al. 1995) and dietary factors. Maintenance of an adequate supply of the ultimate methyl donor (i.e., S-adenosylmethionine) requires an adequate supply of dietary folate and B vitamins. Two previous studies of As metabolism in very small groups of children have suggested that children are poor methylators compared with adults (Chowdhury et al. 2003; Concha et al. 1998). For example, children in two exposed villages in Argentina eliminated 49 and 42%, respectively, as InAs in urine, significantly more than did women in the same villages (25 and 29%, respectively) (Concha et al. 1998). In our study, children were not poor metabolizers. Only 12.2% of urinary As was in the InAs form; mean levels of MMA and DMA were 8.8 and 74.1%, respectively. These metabolite levels compare favorably with those reported for adults (Concha et al. 1998; Hopenhayn-Rich et al. 1996) in various parts of the world and...
with those of a subset of 300 adults in our cohort study (data not shown).

**Mn and Pb exposure.** The relationship between water Mn and children’s intellectual function suggested a possible adverse effect, above and beyond the contribution of social factors. However, that relationship did not persist once water As was added to the regression model. This study was not designed to examine the effects of Mn exposure on intellectual function, and in fact there was a moderate and significant positive association between water Mn and water As. A rigorous examination of the possible relationship between Mn exposure and intellectual function in children calls for a study design in which As exposure is extremely low.

We did not observe the anticipated relationship between BPb and child intellectual function. Our ability to detect this relationship was severely hampered by low statistical power, because approximately half of the study children refused to provide a blood sample.

**Limitations.** We cannot comfortably make a statement about IQ points lost in relationship to As exposure, because of limitations in the application of the U.S. standardization norms to the generation of IQ scores in the present study population. As we have pointed out, the lack of measures of intelligence standardized for use in Bangladesh hampers our ability to draw inferences about IQ points lost at given levels of exposure. Although we have followed sound procedures (derived from our related work in Kosovo) for adapting a widely used instrument to this very different cultural setting, and although we have avoided, for the most part, drawing conclusions about IQ, the measures used here are not measures of IQ, and the absence of standardized measures remains a limitation.

Employing raw scores avoids many pitfalls that would result from using nonstandardized procedures; however, the removal of culturally bound items and subscales diverges from common practice. On the other hand, other simpler predictors of child intellectual function, such as maternal education and child height and head circumference, were significantly related to intellectual raw scores in the expected directions. This gives us confidence in the validity of the observed associations with As. To provide estimates of the impact of As exposure on IQ that would be more directly useful to policy makers, future research should either standardize an IQ test for Bangladesh (a considerable undertaking) or replicate the present effort in a well-defined sample of Western children. Given that the prevalence of malnutrition is quite high in Bangladesh, and that children in our study were of small stature relative to U.S. norms (although not anemic), the dose–response relationship in U.S. children may be different.

The present investigation examines a single age group at a single point in time. We do not know whether the present level of deficit can be detected earlier, whether continued exposure is associated with increased intellectual loss, or, conversely, whether a reduction in exposure would be associated with improved functioning. Better understanding of the exposure–outcome relationship could be obtained by following a group of children from an earlier age and tracking both exposure and outcome regularly.

We believe that our finding of a strong association between As exposure and intelligence is both important and tragic and adds urgency to the need for effective remediation in Bangladesh and other regions of South Asia where consumption of As-contaminated water is prevalent. The global community has been slow in responding to the public health significance of As exposure in Bangladesh, despite the enormous scope of the problem. We hope that the present findings add a new sense of urgency to efforts aimed at alleviating and eliminating As exposure in Bangladesh.

**REFERENCES**


With this in mind, I thank Hopfenberg for his article (2003) in which he provided an elegant model that accounts for the salient factors governing the dynamics of global human population numbers. According to his findings, the size of the human population is determined primarily by food availability.

The realization that these two points of view differ—that there is complexity and simplicity in the world we inhabit—does not necessarily mean that one is correct and the other incorrect. To the contrary, it could be that each point of view is valid based on the scope of observation.

It may be somehow not quite right to agree with the entire idea of Hobbs and Fowler (2004) that “human population size is beyond human capacity to list, comprehend, and synthesize” without noticing that the same can be said regarding any observable phenomenon. Reality is likely just as complex as Hobbs and Fowler described; but it is also clear from the research of Hopfenberg (2003) and Hopfenberg and Pimentel (2001) that the dynamics of human population growth is no longer preternatural but attributable to human cultural processes, consumption of natural resources, and the increasing human population can be seen as patently unsustainable. Understanding the causes of and limits to humanity’s impact in the world is a necessary step toward changing human production, consumption, and population trends. Regardless of how long a culture prizes growth and chooses to leave it unchecked, surely it is not too late to accept limits to growth of the human economy, human consumption, and human numbers worldwide by altering human behavior accordingly.

The author declares he has no competing financial interests.

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Conflicts of Interests:
Declarations for All

Concerning your editorial, “Embracing Scrutiny,” in the October issue of EHP [Environ Health Perspect 112:A788 (2004)], the need for full disclosure of all potential conflicts of interest by all coauthors contributing to a publication in EHP is commendable and obviously needed. Might I take this one step further and suggest that all reviewers of EHP manuscripts be required to sign a form listing all of their potential conflicts of interest.

The author declares he has no competing financial interests.

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Editor’s note: In our Instructions to Authors (http://ehp.niehs.nih.gov/docs/admin/edpolicy.html), we do require editors and reviewers to disclose competing financial interests but are not currently requiring a signed form. We are considering taking that next step.